

## hw5

### 5.1

假设  $\mathbf{x}_{11}, \dots, \mathbf{x}_{1n_1}$  iid  $\sim N_p(\mu_1, \Sigma)$ ,  $\mathbf{x}_{21}, \dots, \mathbf{x}_{2n_2}$  iid  $\sim N_p(\mu_2, \Sigma)$ , 两组样本独立, 两组的样本均值分别是  $\bar{\mathbf{x}}_1, \bar{\mathbf{x}}_2$ , 样本方差矩阵分别是  $S_1, S_2$ 。定义  $S_{\text{pooled}} = [(n_1 - 1)S_1 + (n_2 - 1)S_2] / (n - 2)$ ,  $n = n_1 + n_2$ 。考虑两样本检验问题  $H_0: \mu_1 - \mu_2 = \mathbf{0}_{p \times 1}$ , 两样本 Hotelling  $T^2$  检验统计量定义为

$$T^2 = \frac{n_1 n_2}{n_1 + n_2} (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)^\top S_{\text{pooled}}^{-1} (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2).$$

另一方面, 两样本问题作为多组 MANOVA 的最简单的情况, 其检验也可使用 MANOVA 的 Wilks 检验  $\Lambda^* = |W|/|W + B|$  ( $W, B$  的定义, 第 9 讲 P3), 试验证

$$-n \log \Lambda^* = n \log (1 + T^2 / (n - 2)).$$

#### Solution:

根据定义可知

$$\begin{aligned} W &= \sum_{k=1}^2 (n_k - 1) S_k = (n - 2) S_{\text{pooled}}, \\ B &= \sum_{k=1}^2 n_k (\bar{\mathbf{x}}_k - \bar{\mathbf{x}}) (\bar{\mathbf{x}}_k - \bar{\mathbf{x}})^\top \\ &= \frac{n_1 n_2}{n_1 + n_2} (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2) (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)^\top. \end{aligned}$$

$$\begin{aligned} \Lambda^* &= \frac{|W|}{|W + B|} = \frac{1}{|I + W^{-1}B|} = \left| I + \frac{n_1 n_2}{(n - 2)(n_1 + n_2)} (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2) (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)^\top S_{\text{pooled}}^{-1} \right|^{-1} \\ &= \left( 1 + \frac{n_1 n_2}{(n - 2)(n_1 + n_2)} (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)^\top S_{\text{pooled}}^{-1} (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2) \right)^{-1} \\ &= \left( 1 + \frac{T^2}{n - 2} \right)^{-1}. \end{aligned}$$

所以

$$-n \log \Lambda^* = n \log (1 + T^2 / (n - 2)).$$

■

### 5.2

假设  $\mathbf{x}_1, \dots, \mathbf{x}_n$  iid  $\sim N_p(\mu, \Sigma)$ , 我们考虑如下球对称假设

$$H_0: \Sigma = \gamma I_p, \gamma > 0 \text{ (未知)},$$

记号  $S$  为样本协方差矩阵, 似然比统计量  $\Lambda = \frac{\max L(\mu, \gamma I_p)}{\max L(\mu, \Sigma)}$ 。

(a) 证明 Wilks 统计量

$$\Lambda^* = \Lambda^{2/n} = \frac{\det(S)}{(\text{tr}(S)/p)^p}$$

(b) 已知  $-2 \log(\Lambda) = -n \log(\Lambda^*) \xrightarrow{d} \chi_f^2$ , 求出自由度  $f$ 。

#### Solution:

(a) 似然函数为

$$L(\mu, \Sigma) = \frac{1}{(2\pi)^{np/2} |\Sigma|^{n/2}} \exp \left( -\frac{1}{2} \sum_{i=1}^n (x_i - \mu)^\top \Sigma^{-1} (x_i - \mu) \right).$$

对  $\mu$  和  $\Sigma$  求导可以得到极大似然估计

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n x_i = \bar{x},$$

$$\hat{\Sigma} = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\mu})(x_i - \hat{\mu})^\top = \frac{n-1}{n} S.$$

在原假设  $H_0: \Sigma = \gamma I_p$  下, 对  $\mu$  和  $\gamma$  求导可以得到极大似然估计

$$\hat{\mu}_0 = \bar{x},$$

$$\hat{\gamma}_0 = \frac{1}{np} \sum_{i=1}^n (x_i - \bar{x})(x_i - \bar{x})^\top = \frac{n-1}{np} \text{tr}(S).$$

所以似然比统计量为

$$\Lambda = \frac{\max L(\mu, \gamma I_p)}{\max L(\mu, \Sigma)} = \frac{L(\hat{\mu}_0, \hat{\gamma}_0 I)}{L(\hat{\mu}, \hat{\Sigma})} = \left( \frac{p^p \det(S)}{\text{tr}^p(S)} \right)^{n/2}$$

可得

$$\Lambda^* = \Lambda^{2/n} = \frac{\det(S)}{(\text{tr}(S)/p)^p}$$

(b) 根据 Wilks 定理可得

$$f = \dim(\Theta) - \dim(\Theta_0) = \left( p + \frac{p(p+1)}{2} \right) - (p+1) = \frac{p(p+1)}{2} - 1.$$

### 5.3

假设  $\mathbf{x}_1, \dots, \mathbf{x}_n$  iid  $\sim N_p(\mu, \Sigma)$ ,  $p > 1$ , 样本均值和方差分别是  $\bar{\mathbf{x}}, S$ 。考虑原假设  $H_0: \mu = \mu_0$ ,  $\mu_0$  已知。如果将数据和参数向一维投影方向投影, 利用投影坐标构建一元检验统计量, 效果如何? 对  $\forall \mathbf{a} \in S^{p-1}$ , 考虑

$$H_0(\mathbf{a}): \mathbf{a}^\top \boldsymbol{\mu} = \mathbf{a}^\top \boldsymbol{\mu}_0,$$

因为  $y_1 = \mathbf{a}^\top \mathbf{x}_1, \dots, y_n = \mathbf{a}^\top \mathbf{x}_n$  iid  $\sim N_1(\mathbf{a}^\top \boldsymbol{\mu}, \mathbf{a}^\top \Sigma \mathbf{a})$ ,  $\bar{y} = \mathbf{a}^\top \bar{\mathbf{x}}, s_y^2 = \mathbf{a}^\top S \mathbf{a}$ , 因此  $H_0(\mathbf{a})$  的  $t$  检验统计量

$$t(\mathbf{a}) = \frac{\sqrt{n}(\bar{y} - \mathbf{a}^\top \boldsymbol{\mu}_0)}{s_y} = \frac{\sqrt{n}(\mathbf{a}^\top \bar{\mathbf{x}} - \mathbf{a}^\top \boldsymbol{\mu}_0)}{\sqrt{\mathbf{a}^\top S \mathbf{a}}} \stackrel{H_0(\mathbf{a})}{\sim} t_{n-1}.$$

对所有  $\mathbf{a}$ , 我们取最显著的一个结果, 即取  $t_{\max}^2 = \max_{\mathbf{a} \in S^{p-1}} |t(\mathbf{a})|^2$  作为  $H_0$  的检验统计量。证明

$$t_{\max}^2 = n(\bar{\mathbf{x}} - \boldsymbol{\mu}_0)^\top S^{-1}(\bar{\mathbf{x}} - \boldsymbol{\mu}_0).$$

**Solution:**

$$t^2(\mathbf{a}) = n \frac{\mathbf{a}^\top (\bar{\mathbf{x}} - \boldsymbol{\mu}_0)(\bar{\mathbf{x}} - \boldsymbol{\mu}_0)^\top \mathbf{a}}{\mathbf{a}^\top S \mathbf{a}}$$

令  $u = S^{1/2} \mathbf{a}$ , 根据 Cauchy-Schwarz 不等式

$$\begin{aligned} t^2(\mathbf{a}) &= n \frac{u^\top S^{-\frac{1}{2}} (\bar{\mathbf{x}} - \boldsymbol{\mu}_0)(\bar{\mathbf{x}} - \boldsymbol{\mu}_0)^\top S^{-\frac{1}{2}} u}{u^\top u} \\ &= n \left( \frac{u}{\|u\|} \right)^\top S^{-\frac{1}{2}} (\bar{\mathbf{x}} - \boldsymbol{\mu}_0)(\bar{\mathbf{x}} - \boldsymbol{\mu}_0)^\top S^{-\frac{1}{2}} \left( \frac{u}{\|u\|} \right) \\ &\leq n \left( \frac{u}{\|u\|} \right)^\top \left( \frac{u}{\|u\|} \right) (\bar{\mathbf{x}} - \boldsymbol{\mu}_0)^\top S^{-1} (\bar{\mathbf{x}} - \boldsymbol{\mu}_0) \\ &= n (\bar{\mathbf{x}} - \boldsymbol{\mu}_0)^\top S^{-1} (\bar{\mathbf{x}} - \boldsymbol{\mu}_0) \end{aligned}$$